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Training module # SWDP - 30

# How to validate rating curve

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with HALCROW, TAHAL, CES, ORG & JPS

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While designing a training course, the relationship between this module and the others, would be maintained by keeping them close together in the syllabus and place them in a logical sequence. The actual selection of the topics and the depth of training would, of course, depend on the training needs of the participants, i.e. their knowledge level and skills performance upon the start of the course.

# 2. Module profile

Title	:	How to validate rating curve					
Target group	:	Assistant Hydrologists, Hydrologists, Data Processing Centre Managers					
Duration	:	One session of 60 minutes					
Objectives	:	<ul><li>After the training the participants will be able to:</li><li>Validate the rating curve</li></ul>					
Key concepts	concepts : • Graphical validation • Statistical validation						
Training methods	:	Lecture, software					
Training tools required	:	Board, OHS, Computer					
Handouts	:	As provided in this module					
Further reading : and references							

No	Activities	Time	Tools
1	<i>General</i> Overhead: Text: Rating curve validation (1) Overhead: Text: Rating curve validation (2)	5 min	OHS 1 OHS 2
2	Graphical validation tests Overhead: Figure 2.1: Rating curve fitting Overhead: Table 2.1: Results of curve fitting Overhead: Figure 2.2: Results of validation (1) Overhead: Figure 2.3: Results of validation (2) Overhead: Figure 2.4: Period-flow deviation scatterdiagram Overhead: Figure 2.5: Stage-flow deviation diagram Overhead: Figure 2.6: Cumulative deviation plot	20 min	OHS 3 OHS 4 OHS 5 OHS 6 OHS 7 OHS 8 OHS 9
3	Numerical validation tests Overhead: Text:Use of Student's 't' test to check gaugings Overhead: Table 2.2 Example of 't' test Overhead: Text: Absence from bias test (signs) Overhead: Text: Absence from bias test (values) Overhead: Text: Goodness of fit test	20 min	OHS 10 OHS 11 OHS 12 OHS 13 OHS 14
4	<i>Exercise</i> Validate the gaugings of August 1997 for stations Rakshewa and Khamgaon. Existing rating curve is based on data 1/6- 3/8/97		

Add copy of Main text in chapter 8, for all participants.

# 6. Additional handout

These handouts are distributed during delivery and contain test questions, answers to questions, special worksheets, optional information, and other matters you would not like to be seen in the regular handouts.

It is a good practice to pre-punch these additional handouts, so the participants can easily insert them in the main handout folder.

# 7. Main text

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## 1. General

• Validation of a rating curve is required both after the relationship has first been fitted and subsequently when new gaugings have been carried out, to assess whether these indicate a change in rating. Validation is also used to assess the reliability of historical ratings.

Current meter gauging is carried out with variable frequency depending on previous experience of the stability of the control and of the rating curve. As a minimum it is recommended that six gaugings per year are carried out even with a station with a stable section and previously gauged over the full range of level. At unstable sections many more gaugings are required. The deviation of such check gaugings from the previously established relationship is computed and any bias assessed to determine whether they belong to the same population as the previous stage-discharge relationship.

- Graphical and numerical tests are designed to show whether gaugings fit the current relationship equally and without bias over the full range of flow and over the full time period to which it has been applied. If they do not, then a new rating should be developed as described in Module 29, but taking into account the deficiencies noted in validation.
- Validation will be carried out at Divisional offices or at the State Data Processing Centre.

## 2. Graphical validation tests

#### 2.1 General

Graphical tests are often the most effective method of validation. These include the following:

- Stage/discharge plot with the new gaugings
- Period/flow deviation scattergram
- Stage/flow deviation scattergram
- Cumulative deviation plot of gaugings.
- Stage/discharge plots with gaugings distinguished by season

# Judgements based on graphical displays are often indicative rather than prescriptive - a judgement on the part of the data processor is still required.

#### 2.2 Stage/discharge plot with new gaugings

The simplest means of validating the rating curve with respect to subsequent gaugings is to plot the existing rating curve with the new check gaugings. This is shown in the example for Station Pargaon. A rating curve is established for the period 30/6 - 3/8/97, see Figure 2.1. It shows a proper fit of the data to the existing rating curve, of which the numerical results are shown in Table 2.1. New data are available for the period 4-23/8/97. The new data with the existing rating curve are shown in Figure 2.2. From this plot it is observed that the new gaugings do not match with the existing curve. In Figure 2.3 the new gaugings are shown with the rating curve and its the 95% confidence limits (derived as t-times the standard error S<sub>e</sub>). From this plot it can be judged if most check gaugings lie inside the confidence limits

and thus whether they can be judged acceptable with respect to deviation. It is expected that 19 out of 20 observations will lie inside the limits if the standard error is considered at 5% significance level. However, except insofar as one can see whether all the new points lie

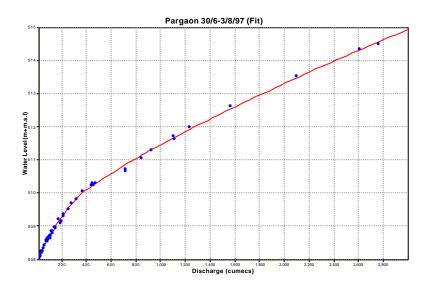


Figure 2.1 Rating curve for station PARGAON established based on data for the period 30/6-3/8/97

#### Table 2.1 Results of rating curve fitting

Analysis of stage-discharge data Station name : PARGAON Data from 1997 6 30 to 1997 8 2 Single channel Given boundaries for computation of rating curve(s) interval lower bound upper bound nr. Of data

49	510.200	508.000	1
12	515.000	509.900	2

Power type of equation q=c\*(h+a)\*\*b is used

		ies / coet ound upper		a	b	С	
			07.730 08.490		1043E+03 1936E+03		
	Number	W level M	Q meas M3/S	Q comp M3/S	DIFf M3/S	Rel.dIFf 0/0	Semr 0/0
	54 56 57 62 63 64 65 66 67	508.050 511.240 510.740 512.620 514.510 514.360 513.530 512.010 511.730	22.830 820.010 711.230 1566.740 2757.370 2609.830 2098.120 1235.470 1103.470	19.116 861.213 640.582 1568.961 2735.225 2635.279 2104.625 1239.485 1096.850	$\begin{array}{r} 3.714 \\ -41.203 \\ 70.648 \\ -2.221 \\ 22.145 \\ -25.448 \\ -6.505 \\ -4.015 \\ 6.620 \end{array}$	19.43 -4.78 11.03 14 .81 97 31 32 .60	4.44 2.58 2.96 3.12 4.60 4.49 3.84 2.71 2.59
	0 11			C 0C1			

Overall standard error = 6.061

Statistics per interval

Interval	Lower bound	. Upper bound	Nr.of data	Standard error
1	508.000	509.981	48	6.55
2	509.981	515.000	12	4.01

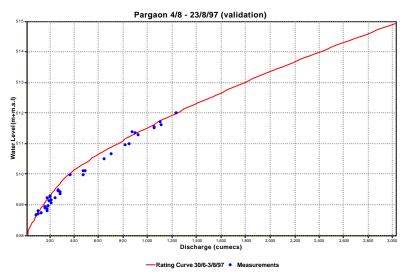
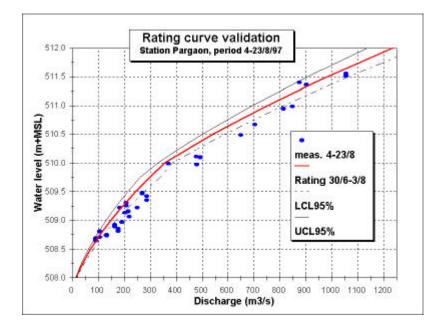


Figure 2.2 New gaugings at Pargaon station plotted against the existing rating curve





above or below the previous regression line, the graph does not specifically address the problem of bias. For example, if some 25 new gaugings may all lie scattered within 95% confidence limits, it does not show any significant change in behaviour. However, if these points are plotted and sequence of each observation is also considered and if upon that a certain pattern of deviation (with respect to time) is perceivable and significant then such situation may warrant new ratings for different periods of distinct behaviour. For the Pargaon-case the plot confirms earlier observations that the new gaugings significantly differ from the existing rating.

#### 2.2 Period/flow deviation scattergram

A period/flow deviation scattergram (Figure 2.4) is a means of illustrating the negative and positive deviation of each current meter gauging from the present rating curve and whether there has been a gradual or sudden shift in the direction of deviations within the period to which the rating has been applied. It also shows whether recent additional gaugings show deviation from previous experience. In the example shown in Fig. 2.4, percentage deviations are very high; there are far more gaugings with positive than with negative deviations. The rating is therefore biased and a revision of the rating is strongly recommended.

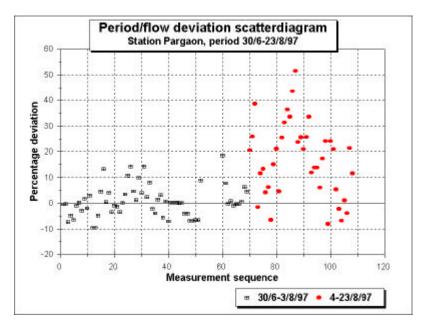
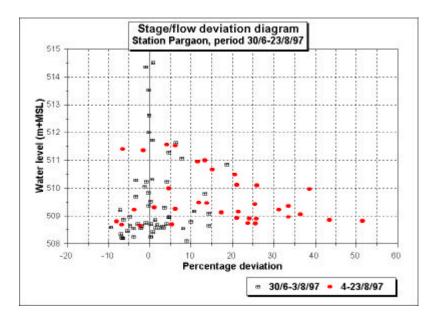


Figure 2.4 Period-flow deviation scatterdiagram for Pargaon rating curve data and new gaugings

### 2.3 Stage/flow deviation diagram

A similar scattergram plot shows the percentage deviation with stage (Figure 2.5) and is a means of illustrating whether over certain ranges of stage the relationship is biased. Most recent gaugings can also be placed within this context.



#### Figure 2.5 Stage-flow deviation scatterdiagram for Pargaon rating curve data and new gaugings

In the example shown in Fig. 2.5, there is some difference in deviation at different stages; particularly at the lower stages the differences are substantial. This plot also confirms the necessity for revision of the rating curve.

### 2.4 Cumulative deviation plot of gaugings

A plot of the cumulative deviation of gaugings from the rating curve give another indication of bias and whether that bias changes with time. Fig. 2.6 shows such a plot for the example of station Pargaon. From the upward trend of the line for the new gaugings it is concluded that the new gaugings produce consistently higher flow values for the same stages than before.

#### 2.5 Stage discharge plots with gaugings distinguished by season.

It is sometimes helpful to separate gaugings between seasons to demonstrate the effect of varying weed growth or other seasonal factors on the stage discharge relationship. The effects of weed growth may be expected to be at a maximum in low flows before the onset of the monsoon; monsoon high flows wash out the weed which increases progressively from the end of the rains. The discharge for given level may thus differ from one month to another. This shows up more clearly in rivers where winter low flows are little affected by weed growth than summer low flows and thus show much smaller spread. Where an auxiliary gauge is available, a backwater rating curve (normal fall method) may be used. Otherwise a simple rating curve may be used in weed absent periods and Stout's shift method during periods of variability.

### 3. Numerical validation tests

#### 3.1 Use of Student's 't' test to check gaugings

A test such as Student's "t" test may be used to decide whether check gaugings can be accepted as part of the homogeneous sample of observations making up the existing stage-discharge curve. Such a test will indicate whether or not the stagedischarge relation requires re-calculation or the section requires recalibration.

In this test, the 't' statistic is calculated as the ratio of the mean deviation and the standard error of the difference of the means as:

$$t = \overline{d_1} / S \tag{1}$$

where  $d_1$  is the mean deviation of the new gaugings (in percent) from the existing curve

and

$$S = a \sqrt{(N+N_1)/NN_1}$$
 (2)

- where *N* is the number of gaugings used to derive the existing rating
- and  $N_1$  is the number of new gaugings
  - *a* is given by the following expression:

$$a = \sqrt{\frac{\sum (d^2) + \sum (d_1 - \overline{d_1})^2}{N + N_1 - 2}}$$
(3)

where  $S(d)^2$  is the sum of the squares of the percent differences for the old gaugings from the existing rating.

If this computed value of 't' = d / S is greater than the critical value of t' for (N + N<sub>1</sub> - 2) degrees of freedom at 95% probability level then further action must be considered. Either the development of a new rating or a request to field staff for additional check gaugings. The critical values of Students 't statistic at the 95% confidence level can be obtained from the standard tables available for the Student's 't' distribution. It should be noted that rating changes are more frequent and more noticeable in the low flow range. Review and validation is therefore done with respect to each range and, unless there is evidence to the contrary, unaffected ranges should retain the old rating but with the range limits adjusted for the new intersection.

As an example, the validation of the new gaugings at station Pargaon is shown in Table 3.1. The results of the 't'-test is seen to support earlier observation of significant deviation.

#### Table 3.1 Results of validation

	on stage-di name : PA om 1997 8						
	n type: Pow Boundari	es 000 509.	Parame 981 -507. 000 -508.	730 1.4	189 104.3 175 193.6		
Data use Interval 1 2	6	nate parame or of est. .546 .013		data			
Number	W level M	Q meas M3/S	Q comp M3/S	DIFf F M3/S	Rel.dIFf 0/0	Semr 0/0	
81 82 83 84 85 86 87 88 89 90 90 91 92 93 94	509.990 509.420 509.060 508.960 508.850 508.810 508.740 508.740 508.920 509.350 509.480 509.460	368.450 285.790 247.940 217.610 189.660 177.270 130.970 130.970 163.570 163.570 285.790 268.610 268.610	104.309 135.149 130.107 213.934 239.991	$16.189 \\ 57.947 \\ 59.060 \\ 58.122 \\ 47.696 \\ 53.788 \\ 60.298 \\ 25.107 \\ 26.661 \\ 28.421 \\ 33.463 \\ 71.856 \\ 28.619 \\ 32.695 \\ \end{cases}$	$\begin{array}{r} 4.60\\ 25.43\\ 31.27\\ 36.44\\ 33.60\\ 43.56\\ 51.55\\ 23.72\\ 25.56\\ 21.03\\ 25.72\\ 33.59\\ 11.93\\ 13.86\end{array}$	$15.48 \\ 15.51 \\ 11.81 \\ 10.23 \\ 10.57 \\ 12.38 \\ 13.39 \\ 15.53 \\ 15.88 \\ 11.05 \\ 11.55 \\ 14.11 \\ 16.73 \\ 16.32 \\ 16.32 \\ 10.51 \\ 10.5$	
Overall	standard e	error =	14.729				
Interval 1 2	508.0 509.9	nd Upper b 00 509 81 515	.981 .000	data Stand 27 15	lard error 25.78 13.18		
Results	Results of student T-test on absence of bias						

RESULTS OF SCULET	IL I	-lest on	au	Serice OI	DIas		
Interval Degrees	of	freedom	95%	T-value	Actual	T-value	Result
1		73		1.993		7.673	Reject
2		25		2.060		2.920	Reject

#### 3.2 Test for absence from bias in signs

A well-balanced rating curve must ensure that the number of positive and negative deviations of the observed values from the rating curve is evenly distributed. That is, the difference in number between the two should not be more than can be explained by chance fluctuations. The test is employed to see if the curve has been established in a balanced manner so that the two sets of discharge values, observed and estimated (from the curve), may be reasonably supposed to represent the same population.

This test is performed by counting observed points falling on either side of the curve. If  $Q_i$  is the observed value and  $Q_c$  the estimated value, then the expression,  $Q_i - Q_c$ , should have an equal chance of being positive or negative. In other words, the probability of  $Q_i - Q_c$  being positive or negative is  $\frac{1}{2}$ . Hence, assuming the successive signs to be independent of each other, the sequence of the differences may be considered as distributed according to the binomial law (p+q) N, where N is the number of observations, and p and q, are the probabilities of occurrence of positive and negative values are  $\frac{1}{2}$  each. The expected number of positive signs is positive signs is Np. Its standard deviation is  $\sqrt{(Npq)}$ . The "t" statistic is then found by dividing the difference between the actual number of positive signs N<sub>1</sub> and expected number of positive signs Np by its standard deviation **Ü**(Npq):

$$t = \frac{\left|N_1 - Np\right| - 0.5}{\sqrt{Npq}}$$

The resulting value is compared with the critical value of "t" statistic for 5% significance level for the degrees of freedom equal to the total number of stage discharge data. If the value of the critical "t" statistic is more than that obtained for the observed data then it can be considered that the data does not show any bias with respect to sign of the deviations between observed and computed discharges.

#### 3.3 Test for absence from bias in values

This test is designed to see if a particular stage discharge curve, on average, yields significant under estimates or over estimates as compared to the actual observations on which it is based. (Compare the graphical test using the period/flow deviation and stage /flow deviation scattergrams) The percentage differences are first worked out as:

$$P = 100 (Q_i - Q_c) / Q_c$$
 (4)

If there are N observations and  $P_1$ ,  $P_2$ ,  $P_3$ , ...,  $P_N$  are the percentage differences and  $P_{av}$  is the average of these differences, the standard error of  $P_{av}$  is given by:

$$Se = \sqrt{\frac{(\sum P_{i} - P_{av})^{2}}{N(N-1)}}$$
(5)

The average percent  $P_{av}$  is tested against its standard error to see if it is significantly different from zero. The "t" statistic for in this case is computed as:

$$t = (P_{av} - 0) / S_e$$
 (6)

If the critical value of "t" statistic for 5% significance level and N degrees of freedom is greater than the value computed above then it may be considered that there is no statistical bias in the observed values with respect to their magnitudes as compared with that obtained by the rating curve.

The percentage differences have been taken as they are comparatively independent of the discharge volume and are approximately normally distributed about zero mean value for an unbiased curve.

#### 3.4 Goodness of fit test

Due to changes in the flow regime, it is possible that long runs of positive and/or negative deviations are obtained at various stages. This may also be due to inappropriate fitting of the rating curve. This test is carried out for long runs of positive and negative deviations of the observed values from the stage-discharge curve. The test is designed to ensure a balanced fit in reference to the deviations over different stages. (Compare the graphical tests using stage/flow deviation scattergram and cumulative deviation plot of gaugings)

The test is based on the number of changes of sign in the series of deviations (observed value minus expected or computed value). First of all, the signs of deviations, positive or negative, in discharge measurements in ascending order of stage are recorded. Then starting from the second sign of the series, "0" or "1" is placed under sign if the sign agrees or does not agree respectively with the sign immediately preceding it. For example,

+-++++ -- -- +++++-11000100010001

If there are N numbers in the original series, then there will be (N - 1) numbers in the derived series 11000100010001

If the observed values are regarded as arising from random fluctuations about the values estimated from the curve, the probability of a change in sign could be taken to be  $\frac{1}{2}$ . However, this assumes that the estimated value is the median rather than the mean. If N is fairly large, a practical criterion may be obtained by assuming successive signs to be independent (i.e. by assuming that they arise only from random fluctuations), so that the number of "1"s (or "0"s) in the derived sequence of (N - 1) members may be judged as a binomial variable with parameters (N - 1) and  $\frac{1}{2}$ .

From the above derived series, the actual number of changes of sign is noted. The expected number of changes of sign is computed by multiplying total possible numbers (i.e. N-1) with the probability of change of sign (i.e.  $\frac{1}{2}$ ). The statistical significance of the departure of the actual number of change of signs from the expected number is known by finding the "t" statistic as follows:

$$t = \frac{|N' - (N-1)p| - 0.5}{\sqrt{(N-1)pq}}$$
(7)

where N' denotes the actual number changes of sign.

If the critical value of "t" statistic, for (N - 1) degrees of freedom, is more than that computed above then it can be considered to be having adequate goodness of fit. Otherwise, the results will indicate that there is significant bias in the fitted curve with respect to long runs of positive or negative deviations.